

Sampling with replacement: Lecture 3

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number of **all outcomes**:

5

five options on the first day

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number of **all outcomes**:



on the first two days $5 \cdot 5 = 25$ options
each cup goes with any other, we can take a product

the cups are different

all outcomes are equally likely

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number of **all outcomes**:



on the first three days $5 \cdot 5 \cdot 5 = 125$ options

each cup goes with any other, we can take a product

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number of **all outcomes**:



there are 5^7 equally likely options during a week
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$$3 \cdot 3 \cdot 3 \cdot 3 \cdot 2 \cdot 2 \cdot 2 = 3^4 \cdot 2^3 \text{ such options}$$

each cup goes with any other, we can multiply

the cups are different

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number of good outcomes:

number of colorings

number of outcomes with a given coloring

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number of good outcomes:

number of colorings \times number of outcomes with a given coloring

$$\binom{7}{4} \times 3^4 \cdot 2^3 = 35 \cdot 3^4 \cdot 2^3$$

the number of ways to choose four days out of seven
the same for each coloring

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$$\mathbb{P}(\text{blue cup exactly four times}) = \frac{\binom{7}{4} \cdot 3^4 \cdot 2^3}{5^7},$$

that is, $\frac{35 \cdot 81 \cdot 8}{78125} = 29,03\%$.

Sampling with replacement

- There are N balls in a box, among which M is red, the others are blue.
- **By putting the balls back after each draw**, we choose n balls, forgetting all previous events, and choosing all N balls with the same probability.
- The probability that there are k red balls among the n chosen ones:

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which red balls were chosen (k draws)

which k choices are red which blue ones were chosen ($n - k$ draws)

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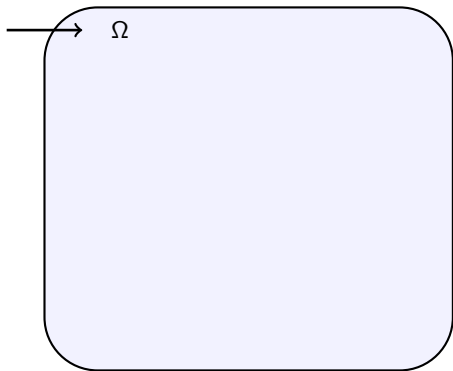
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Remark: the number of chosen red balls has **binomial distribution**.
Since all N^n cases were uniformly likely, this is also **a classical probability space**.

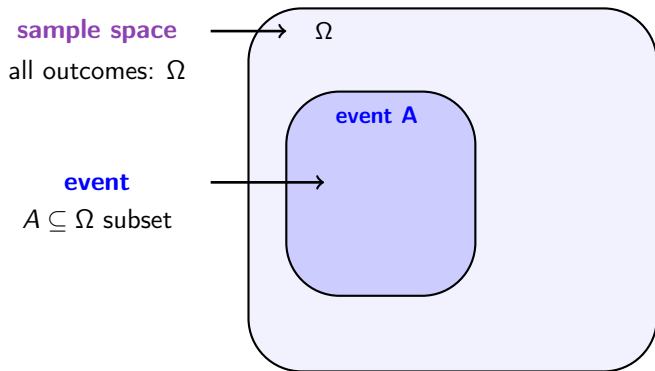
Kolmogorov probability space

sample space

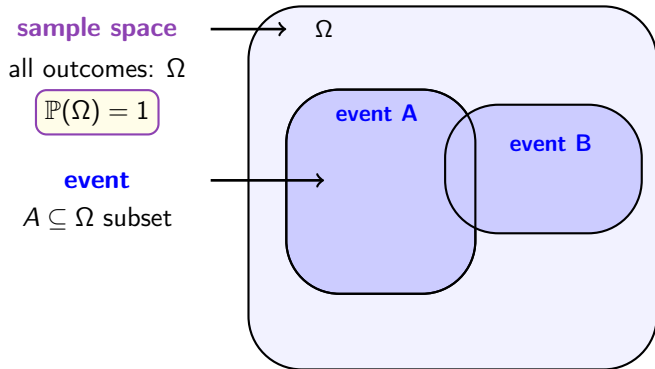
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Ω : (Hungarian) **adult people**

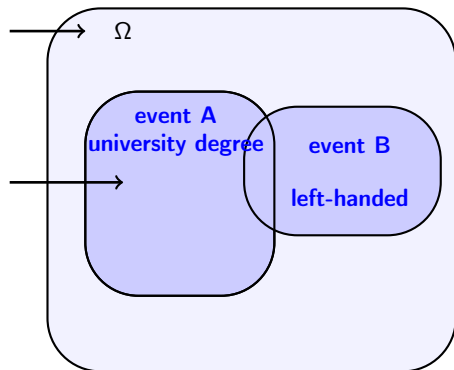
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$$\mathbb{P}(\Omega) = 1$$

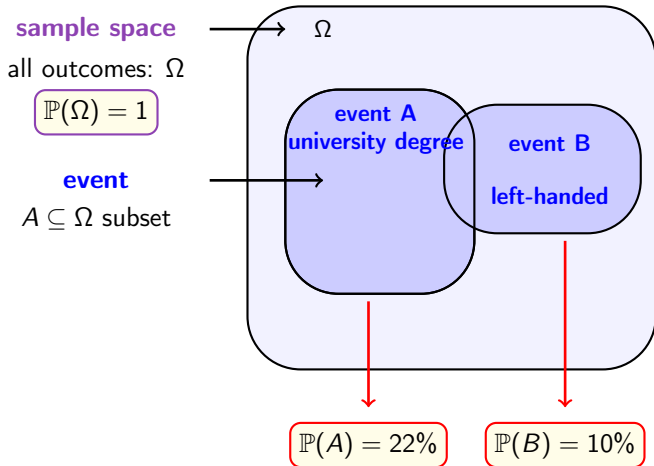
event

$A \subseteq \Omega$ subset



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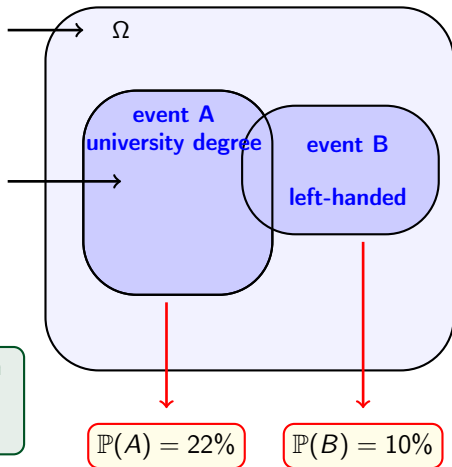
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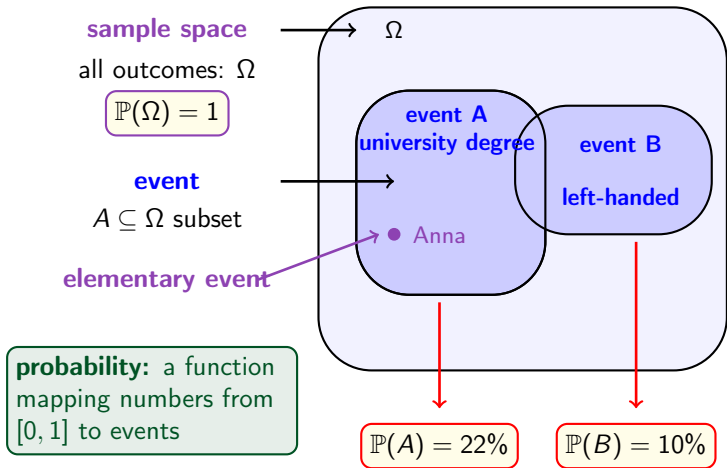
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probability: a function mapping numbers from $[0, 1]$ to events



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Basic properties of probability

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The same with another notation:

$$\mathbb{P}\left(\bigcup_{j=1}^{\infty} A_j\right) = \sum_{j=1}^{\infty} \mathbb{P}(A_j).$$

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while the **impossible event** (empty set) has probability 0.

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- **Probability of complement:** if $A \subseteq \Omega$ is an event, then the probability that **A does not occur:**

$$\mathbb{P}(\bar{A}) = \mathbb{P}(\Omega \setminus A) = 1 - \mathbb{P}(A).$$

- **Probability of difference:**

$$\mathbb{P}(A \setminus B) = \mathbb{P}(A) - \mathbb{P}(A \cap B).$$

(Kolmogorov) probability space

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 - i) $\Omega \in \mathcal{A}$;
 - ii) if $A_1, A_2, \dots \in \mathcal{A}$, then $\bigcup_{n=1}^{\infty} A_n \in \mathcal{A}$ (a countable union of events is also an event);
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- **probability** is a function $\mathbb{P} : \mathcal{A} \rightarrow [0, 1]$, such that
 - ❶ $\mathbb{P}(\Omega) = 1$, that is, the sample space has probability 1;
 - ❷ if $A_1, A_2, \dots \in \mathcal{A}$ and for every $1 \leq i < j$ we have $A_i \cap A_j = \emptyset$, then

$$\mathbb{P}\left(\bigcup_{n=1}^{\infty} A_n\right) = \sum_{n=1}^{\infty} \mathbb{P}(A_n),$$

that is, the probability of the union of countably many, pairwise disjoint events is the sum of the probabilities.

Finite probability space

Suppose that there are finitely many outcomes: $\Omega = \{\omega_1, \omega_2, \dots, \omega_n\}$, and \mathcal{A} is the set of all subsets of Ω .

Notation: $p_j = \mathbb{P}(\{\omega_j\})$ is the probability of the j th outcome. By additivity, we have

$$1 = \mathbb{P}(\Omega) = p_1 + p_2 + \dots + p_n = \sum_{j=1}^n p_j,$$

that is, the probability of the elementary events sums up to 1. Furthermore,

$$\mathbb{P}(A) = \mathbb{P}\left(\bigcup_{j: \omega_j \in A} \{\omega_j\}\right) = \sum_{j: \omega_j \in A} p_j,$$

meaning that

the probability of an event is the sum of the probabilities corresponding to the elementary events that it consists of.

Classical probability space

Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space such that

- Ω is a finite set;
- \mathcal{A} is the set of all subsets of Ω ;
- **all elementary events are equally likely**, that is,

$$\mathbb{P}(\omega_j) = p_j = \frac{1}{n} \quad \text{for all } j = 1, 2, \dots, n.$$

Then $(\Omega, \mathcal{A}, \mathbb{P})$ is a **classical probability space**. In this case, for every event $A \in \mathcal{A}$ we have

$$\mathbb{P}(A) = \frac{k}{n} = \frac{1}{n} + \frac{1}{n} + \dots + \frac{1}{n},$$

where k is the size of A , and n is the total number of elementary events (outcomes).

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Examples: sampling with and without replacement, in both cases, elementary events were equally likely

Example: dice rolls

We take two fair dice rolls. What is the probability that the sum is 7?

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Dice, people, objects etc. are always different.

- **sample space:** possible sequences. Their number:

$$6 \cdot 6 = 36; \text{ both value can be chosen } 6 \text{ ways.}$$

- The sequences are equally likely: each has probability $1/36$.
- Number of good sequences: 6.

11	12	13	14	15	16
21	22	23	24	25	26
31	32	33	34	35	36
41	42	43	44	45	46
51	52	53	54	55	56
61	62	63	64	65	66

Hence $\mathbb{P}(\text{the sum is } 7) = 6/36 = 1/6$.

Counting and notation

$$n! = n \cdot (n-1) \cdot (n-2) \cdot \dots \cdot 2 \cdot 1; \quad \binom{n}{k} = \frac{n!}{k!(n-k)!} = \frac{n(n-1)\dots(n-k+1)}{k(k-1)\dots 2 \cdot 1}.$$

- n objects have $n!$ different orders.
- from n objects, we can choose k different $n(n-1)\dots(n-k+1)$ many ways, if order matters: n ways for the first, $n-1$ for the second (independently of the first one), $n-2$ for the third (independently of the first two), and so on
- from n object a group of size k can be choosen $\binom{n}{k}$ many ways (order does not matter here)
- if for n experiments, we have k options for each, then the total number of outcomes (with order) is

$$n \cdot n \cdot \dots \cdot n = n^k.$$

For example, two dice rolls have 36 possibilities, three rolls have $6 \cdot 6 \cdot 6$, and n have 6^n .

Properties

$$\binom{n}{k} = \binom{n}{n-k}; \quad \binom{n}{0} = 1; \quad \binom{n}{1} = n; \quad \binom{n}{2} = \frac{n(n-1)}{2}$$
$$\binom{n}{N} = 0, \text{ if } N > n.$$

$(x + y)^2 = x^2 + 2xy + y^2$ is generalized to the **binomial theorem**:

$$(x + y)^n = x^n + nx^{n-1}y + \binom{n}{2}x^{n-2}y^2 + \dots + \binom{n}{k}x^k y^{n-k} + \dots + nxy^{n-1} + y^n.$$

Consequence for $x = y = 1$:

$$2^n = \binom{n}{0} + \binom{n}{1} + \binom{n}{2} + \dots + \binom{n}{k} + \dots + \binom{n}{n-1} + \binom{n}{n}.$$

Multinomial/polynomial theorem

- $(x + y)^2 = x^2 + 2xy + y^2$
- $(x + y + z)^2 = x^2 + y^2 + z^2 + 2xy + 2xz + 2yz$
- $(x + y)^3 = x^3 + 3x^2y + 3xy^2 + y^3$

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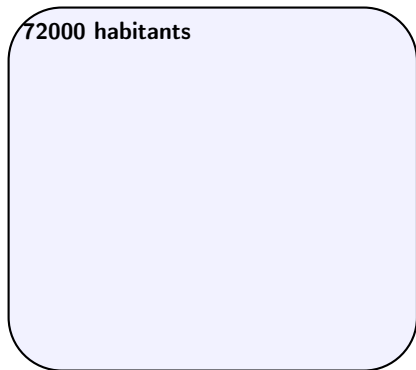
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- $(x + y)^3 = x^3 + 3x^2y + 3xy^2 + y^3$
- in general:

$$(x_1 + x_2 + \dots + x_k)^n = \sum \frac{n!}{i_1! \cdot i_2! \cdot \dots \cdot i_k!} \cdot x_1^{i_1} \cdot x_2^{i_2} \cdot \dots \cdot x_k^{i_k},$$

where we sum up to all sequences of positive integers (i_1, i_2, \dots, i_k) , for which $i_1 + i_2 + \dots + i_k = n$.

Probability of the union: example

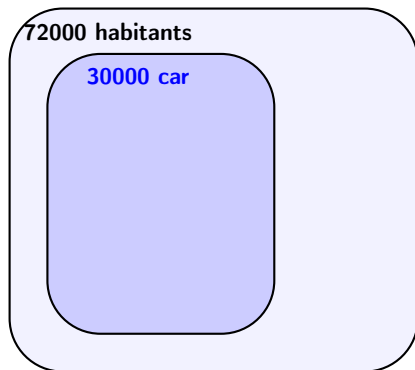
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Probability of the union: example

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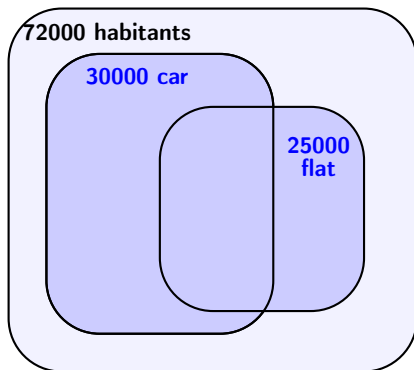
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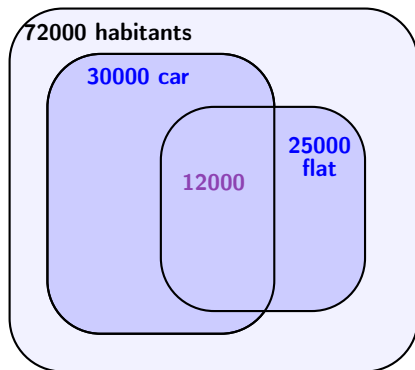
In a city with **72000** habitants
25000 have a flat,

30000 have a car,



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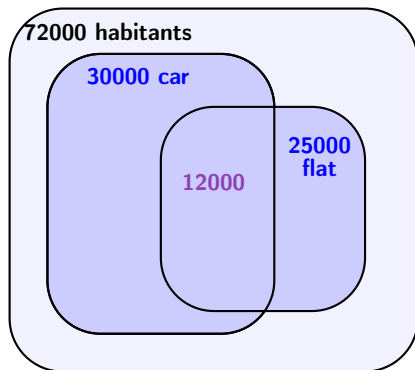
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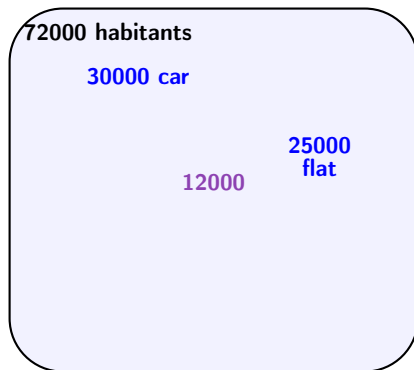
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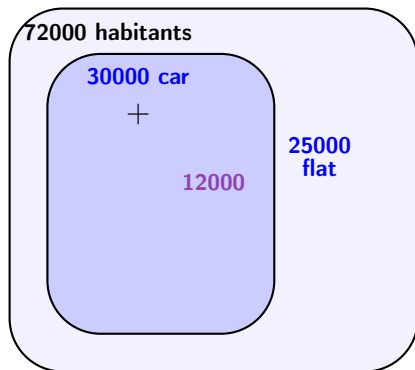
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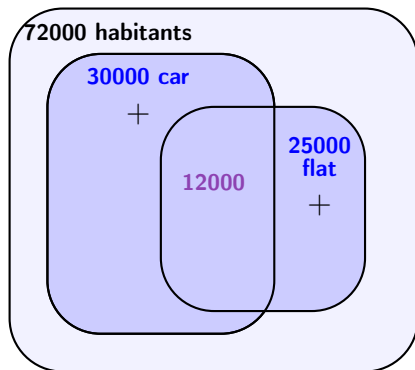
$$\mathbb{P}(\text{at least one}) = \mathbb{P}(\text{car})$$



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$$\mathbb{P}(\text{at least one}) = \mathbb{P}(\text{car}) + \mathbb{P}(\text{flat})$$

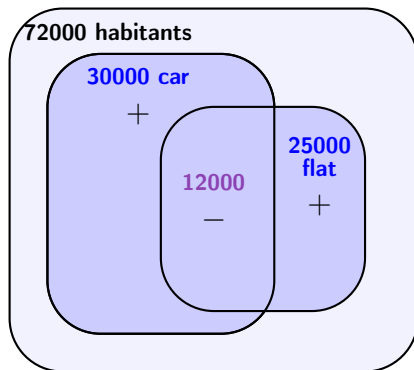


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What is the probability that a uniformly randomly chosen habitant has **at least one of** a car and a flat?

$$\mathbb{P}(\text{at least one}) = \mathbb{P}(\text{car}) + \mathbb{P}(\text{flat}) - \mathbb{P}(\text{both})$$

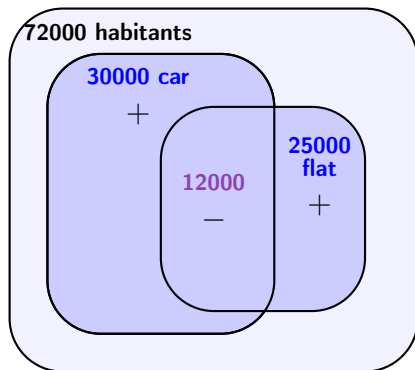
12000 habitants were counted twice, we subtract this



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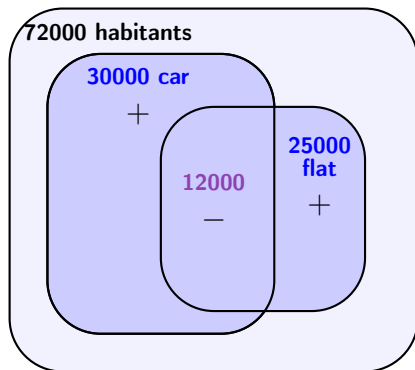
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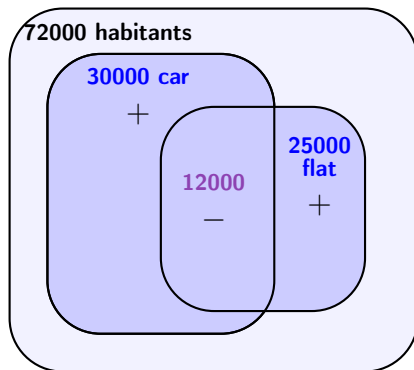


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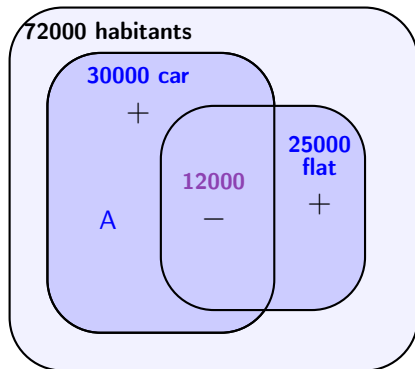


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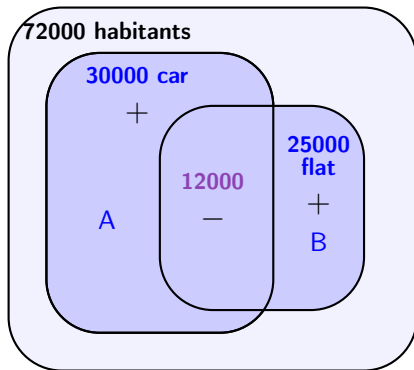


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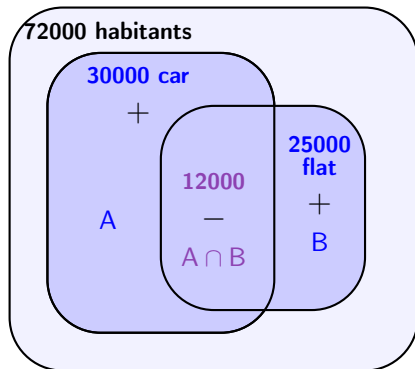
Probability of the union: example

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↑
union

↑
intersection

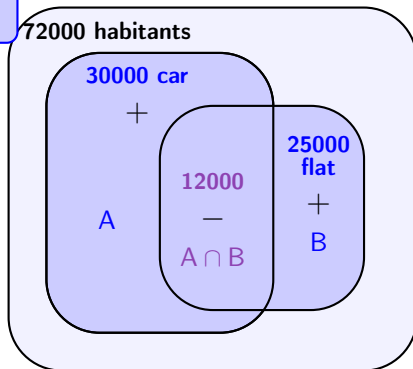


Probability of the union: example

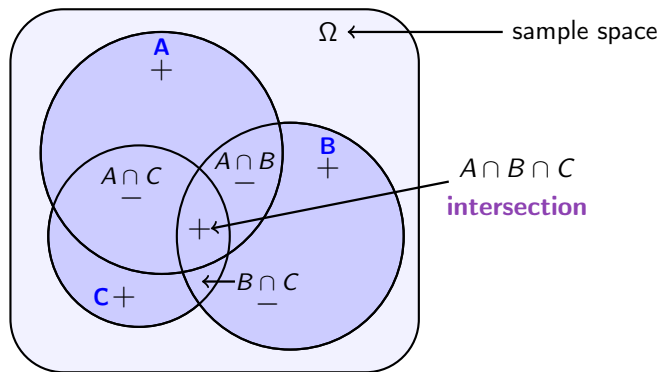
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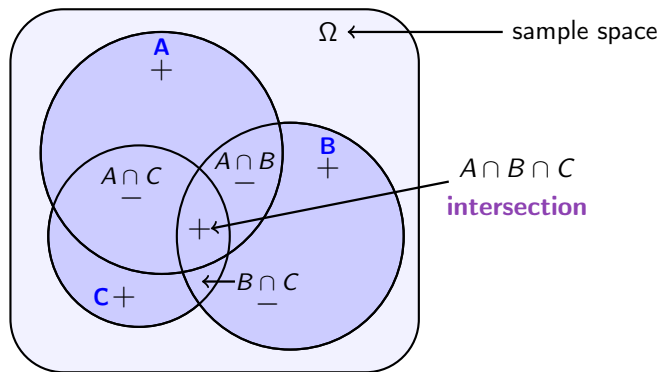
inclusion-
exclusion
formula



Inclusion-exclusion formula for three events



Inclusion-exclusion formula for three events



The probability of the **union** (at least one occurs):

$$\begin{aligned} \mathbb{P}(A \cup B \cup C) &= \mathbb{P}(A) + \mathbb{P}(B) + \mathbb{P}(C) \\ &\quad - \mathbb{P}(A \cap B) - \mathbb{P}(A \cap C) - \mathbb{P}(B \cap C) + \\ &\quad + \mathbb{P}(A \cap B \cap C). \end{aligned}$$

Inclusion-exclusion formula

For two events. The probability that at least one of A and B occurs:

$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B).$$

For three events. The probability that at least one of A , B and C occurs:

$$\begin{aligned} \mathbb{P}(A \cup B \cup C) = & \mathbb{P}(A) + \mathbb{P}(B) + \mathbb{P}(C) - \mathbb{P}(A \cap B) - \\ & - \mathbb{P}(A \cap C) - \mathbb{P}(B \cap C) + \mathbb{P}(A \cap B \cap C) \end{aligned}$$

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In general: the union of A_1, \dots, A_n (at least one event occurs) has probability

$$\begin{aligned} \mathbb{P}\left(\bigcup_{i=1}^n A_i\right) = & \mathbb{P}(A_1) + \mathbb{P}(A_2) + \dots + \mathbb{P}(A_n) - \mathbb{P}(A_1 \cap A_2) - \mathbb{P}(A_1 \cap A_3) - \\ & - \dots - \mathbb{P}(A_{n-1} \cap A_n) + \mathbb{P}(A_1 \cap A_2 \cap A_3) + \dots \end{aligned}$$

Inclusion-exclusion formula

The probability that at least one of the events A_1, A_2, \dots, A_n occurs:

$$\begin{aligned}\mathbb{P}\left(\bigcup_{i=1}^n A_i\right) &= \mathbb{P}(A_1) + \mathbb{P}(A_2) + \dots + \mathbb{P}(A_n) - \mathbb{P}(A_1 \cap A_2) - \mathbb{P}(A_1 \cap A_3) - \\ &\quad - \dots - \mathbb{P}(A_{n-1} \cap A_n) + \mathbb{P}(A_1 \cap A_2 \cap A_3) + \dots\end{aligned}$$

That is:

$$\begin{aligned}\mathbb{P}\left(\bigcup_{i=1}^n A_i\right) &= \sum_{i=1}^n \mathbb{P}(A_i) - \sum_{1 \leq i_1 < i_2 \leq n} \mathbb{P}(A_{i_1} \cap A_{i_2}) + \\ &\quad + \sum_{1 \leq i_1 < i_2 < i_3 \leq n} \mathbb{P}(A_{i_1} \cap A_{i_2} \cap A_{i_3}) - \\ &\quad - \dots + (-1)^{n+1} \mathbb{P}(A_1 \cap \dots \cap A_n) = \\ &= \sum_{k=1}^n (-1)^{k+1} \sum_{1 \leq i_1 < i_2 < \dots < i_k \leq n} \mathbb{P}(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_k}).\end{aligned}$$

Inclusion-exclusion formula: example

Peter collects toys that are sold in a chocolate egg. There are ten types of toys, each egg contains one uniformly at random. What is the probability that his collection is complete (contains all ten types) after buying 20 chocolate eggs?

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First term of the inclusion-exclusion formula:

$$\mathbb{P}(A_1) + \mathbb{P}(A_2) + \dots + \mathbb{P}(A_{10}) = 10 \cdot \frac{9^{20}}{10^{20}},$$

because A_j means that j cannot occur, but all the other 9 can be chosen.

Inclusion-exclusion formula: example

Similarly, the next sum, with a minus sign:

$$\mathbb{P}(A_1 \cap A_2) + \mathbb{P}(A_1 \cap A_3) + \dots + \mathbb{P}(A_9 \cap A_{10}) = \binom{10}{2} \cdot \frac{8^{20}}{10^{20}},$$

because the number of terms is the same as the pairs from $1, 2, \dots, 10$, and $A_i \cap A_j$ means that i and j cannot occur, so we can choose from 8 types for every chocolate egg.

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Altogether:

$$\begin{aligned} \mathbb{P}(A_1 \cup \dots \cup A_{10}) &= 10 \cdot \frac{9^{20}}{10^{20}} - \binom{10}{2} \cdot \frac{8^{20}}{10^{20}} + \binom{10}{3} \cdot \frac{7^{20}}{10^{20}} - \dots = \\ &= \sum_{k=1}^{10} (-1)^{k+1} \binom{10}{k} \frac{(10-k)^{20}}{10^{20}} = 78,53\%. \end{aligned}$$

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This was the probability of the event that at least one type is missing, and hence the probability that the collection is complete is $100 - 78,53 = 21,47\%$.

Independence

Which events can be considered independent? Anne is a randomly chosen participant for a survey.

Anne has a car

rain in Buda tomorrow

Anne's income is more than the average

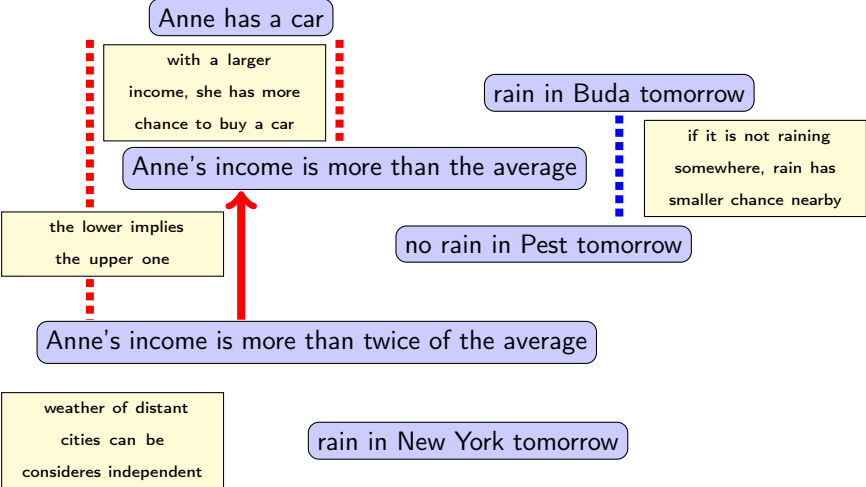
no rain in Pest tomorrow

Anne's income is more than twice of the average

rain in New York tomorrow

Independence

Which events can be considered independent? Anne is a randomly chosen participant for a survey.



Independent events

Events $A, B \in \mathcal{A}$ are **independent**, if

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B),$$

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that is, **the probability of the intersection is the product of the probabilities.**

For several events, all subsets have to satisfy this property.

Events $A_1, A_2, \dots \in \mathcal{A}$ are **independent**, if for every $k \geq 1$ and $1 \leq i_1 < i_2 < \dots < i_k \leq n$ we have

$$\mathbb{P}(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_k}) = \mathbb{P}(A_{i_1})\mathbb{P}(A_{i_2}) \dots \mathbb{P}(A_{i_k}).$$

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31	32	33	34	35	36
41	42	43	44	45	46
51	52	53	54	55	56
61	62	63	64	65	66

A: first roll is 6; **B**: second roll is 6; **$A \cap B$** : both rolls are 6

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A: first roll is 6; **B**: second roll is 6; **$A \cap B$** : both rolls are 6

$$\frac{1}{36} = \mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B) = \frac{1}{6} \cdot \frac{1}{6}.$$

Events A and B are **independent**.

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$$\frac{1}{36} = \mathbb{P}(A \cap C) \neq \mathbb{P}(A) \cdot \mathbb{P}(C) = \frac{1}{6} \cdot \frac{1}{12}.$$

Events A and C are **not independent**.

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41	42	43	44	45	46
51	52	53	54	55	56
61	62	63	64	65	66

A: first roll is 6; **D**: second roll is 7; **$A \cap D$** : first roll is 6, second is 1

Independence: example

We take two fair dice rolls. Are the following events independent: **first roll is 6;**
the sum is 7?

11	12	13	14	15	16
21	22	23	24	25	26
31	32	33	34	35	36
41	42	43	44	45	46
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A: first roll is 6; **D**: second roll is 7; **A ∩ D**: first roll is 6, second is 1

$$\frac{1}{36} = \mathbb{P}(A \cap D) = \mathbb{P}(A) \cdot \mathbb{P}(D) = \frac{1}{6} \cdot \frac{1}{6}.$$

Events A and D are **independent**.